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Research Article

Optimal Bounds for the Variance of Self-Intersection Local Times

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For a \mathbb{Z}^d -valued random walk $(S_n)_{n \in \mathbb{N}_0}$, let $l(n, x)$ be its local time at the site $x \in \mathbb{Z}^d$. For $\alpha \in \mathbb{N}$, define the α -fold self-intersection local time as $L_n(\alpha) := \sum_x l(n, x)^\alpha$. Also let $L_n^{\text{SRW}}(\alpha)$ be the corresponding quantities for the simple random walk in \mathbb{Z}^d . Without imposing any moment conditions, we show that the variance of the self-intersection local time of any genuinely d -dimensional random walk is bounded above by the corresponding quantity for the simple symmetric random walk; that is, $\text{var}(L_n(\alpha)) = O(\text{var}(L_n^{\text{SRW}}(\alpha)))$. In particular, for any genuinely d -dimensional random walk, with $d \geq 4$, we have $\text{var}(L_n(\alpha)) = O(n)$. On the other hand, in dimensions $d \leq 3$ we show that if the behaviour resembles that of simple random walk, in the sense that $\liminf_{n \rightarrow \infty} \text{var}(L_n(\alpha)) / \text{var}(L_n^{\text{SRW}}(\alpha)) > 0$, then the increments of the random walk must have zero mean and finite second moment.

1. Introduction and Main Results

Let X, X_1, X_2, \dots be independent, identically distributed, \mathbb{Z}^d -valued random variables, and define the random walk $S_0 := 0$, $S_n = \sum_{j=1}^n X_j$, for $n \geq 1$. The special case with $\mathbb{P}(X_i = e) = 1/(2d)$, for all $e \in \mathbb{Z}^d$ with $|e| = 1$, is known as the *simple random walk* in \mathbb{Z}^d and will be denoted by $(\text{SRW}_n)_{n \in \mathbb{N}_0}$.

Let $l(n, x) = \sum_{j=1}^n \mathbb{1}(S_j = x)$ be the local time of $(S_n)_{n \in \mathbb{N}_0}$ at the site $x \in \mathbb{Z}^d$, and define for a positive integer α the α -fold *self-intersection local time*

$$\begin{aligned} L_n = L_n(\alpha) &= \sum_{x \in \mathbb{Z}^d} l(n, x)^\alpha \\ &= \sum_{i_1, \dots, i_\alpha=0}^n \mathbb{1}(S_{i_1} = \dots = S_{i_\alpha}). \end{aligned} \quad (1)$$

We will denote the corresponding quantities for simple random walk in \mathbb{Z}^d by $L_n^{\text{SRW}}(\alpha, d)$ or simply $L_n^{\text{SRW}}(\alpha)$ when the dimension is clear from the context.

Let R^+ and R^- be, respectively, the semigroup and the group generated by the support of X ,

$$\begin{aligned} R^+ &:= \{x \in \mathbb{Z}^d \mid \mathbb{P}(S_n = x) > 0 \text{ for some } n \geq 0\}, \\ \bar{R} &:= \{x \in \mathbb{Z}^d \mid x = y - z \text{ for some } x, y \in R^+\}. \end{aligned} \quad (2)$$

Following Spitzer [1], we call the random variable X and the random walk it generates *genuinely d -dimensional* if the group \bar{R} is d -dimensional.

The quantity $L_n(\alpha)$ has received considerable attention in the literature due to its relation to *self-avoiding walks* and *random walks in random scenery*. In particular let the *random scenery* $\{\xi_x, x \in \mathbb{Z}^d\}$ be a collection of i.i.d. random variables, independent of $(S_n)_n$, and define the process $Z_0 = 0$, $Z_n = \sum_{i=1}^n \xi_{S_i}$. Then $(Z_n)_n$ is commonly referred to as *random walk in random scenery* and was introduced in Kesten and Spitzer [2], where functional limit theorems were obtained for $Z_{[nt]}$ under appropriate normalization for the case $d = 1$. The case $d = 2$, with X_i centered with nonsingular covariance matrix, was treated in [3] where it

was shown that $Z_{[nt]}/\sqrt{n \log n}$ converges weakly to Brownian motion. As is obvious from the identities $Z_n = \sum_{x \in \mathbb{Z}^d} l(n, x) \xi_x$ and $\text{var}(Z_n) = \text{var}[L_n(2)] \text{var}(\xi_x)$, limit theorems for $(Z_n)_n$ usually require asymptotic results for the local times of the random walk $(S_n)_n$.

Such asymptotic results are usually obtained from Fourier techniques applied to the characteristic function $f(t) = \mathbb{E}[\exp(it \cdot X)]$, under the additional assumption of a Taylor expansion of the form $f(t) = 1 - \langle \Sigma t, t \rangle + o(|t|^2)$, where Σ is a positive definite covariance matrix [3–7], which further requires that $\mathbb{E}|X|^2 < \infty$ and $\mathbb{E}X = 0$. Similar restrictions are also required for the application of local limit theorems such as in [8, 9].

In this paper, motivated by the results of Spitzer [1] for genuinely d -dimensional random walks and the approach of Becker and König [10], we will study the asymptotic behavior of $\text{var}(L_n(\alpha))$ without imposing any moment assumptions on the random walk. The central idea behind our approach is to compare the self-intersection local times $L_n(\alpha)$ of a general d -dimensional walk with those of its symmetrised version. In addition we will compare the self-intersection local times of a general d -dimensional random walk with those of the d -dimensional simple symmetric random walk, $(\text{SRW}_n)_{n \in \mathbb{N}_0}$. It is well known that, for some positive constants $K_{\alpha, d}$, $\text{var}(L_n^{\text{SRW}}(\alpha, d)) \sim K_{\alpha, d} v_{d, \alpha}(n)$ as $n \rightarrow \infty$, for

$$\begin{aligned} v_{1, \alpha}(n) &:= n^{1+\alpha}, \\ v_{2, \alpha}(n) &:= n^2 \log(n)^{2\alpha-4}, \\ v_{3, \alpha}(n) &:= n \log(n), \\ v_{d, \alpha}(n) &:= n, \quad d \geq 4. \end{aligned} \quad (3)$$

Several other cases have been treated in the literature, using a variety of methods.

A careful look at the literature reveals that the most difficult case in $d = 2$ is the *near transient recurrent* case, where $\mathbb{P}(S_n = 0) \sim C/n$, which corresponds to genuinely 2-dimensional symmetric recurrent random walks, which will be referred to as a critical case. Surprisingly enough, the variance of the self-intersection local times in the critical case is asymptotically the largest.

Theorem 1. *Let X, X_1, X_2, \dots be independent, identically distributed, and genuinely d -dimensional \mathbb{Z}^d -valued random variables, for any $d \geq 1$. Then, there exist positive constants $C_{\alpha, X} > c_{\alpha, X} > 0$, depending on α and the distribution of X , such that for all n large enough*

$$\text{var}(L_n(\alpha)) \leq c_{\alpha, X} \text{var}(L_n^{\text{SRW}}(\alpha, d)) \leq C_{\alpha, X} v_{d, \alpha}(n). \quad (4)$$

The result was motivated by [1, 10] and improves related results of Becker and König for $d = 3$ and $d = 4$. Several cases treated in [3, 4, 10–13] can then be obtained as particular cases.

Moreover, we also show the surprising converse. More precisely, we show that the right asymptotic behaviour of $\text{var}(L_n)$ implies that the jumps must have zero mean and finite second moment.

Theorem 2. *Let X, X_1, X_2, \dots be independent, identically distributed, and genuinely d -dimensional with $d \leq 3$. If*

$$\liminf_{n \rightarrow \infty} \frac{\text{var}(L_n(\alpha))}{\text{var}(L_n^{\text{SRW}}(\alpha))} > 0, \quad (5)$$

then $\mathbb{E}|X|^2 < \infty$ and $\mathbb{E}X = 0$.

As it follows from Theorem 3 given below for $d = 2, 3$ and from Theorem 5.2.3 in Chen [12] for $d = 1$, if $\mathbb{E}X = 0$ and $0 < \mathbb{E}|X|^2 < \infty$, then $\liminf_n \text{var}(L_n(\alpha))/v_{d, \alpha}(n) > 0$.

For any genuinely d -dimensional random walk with finite second moments and zero mean, the asymptotic behaviour of $\text{var}(L_n(\alpha))$ is similar to that of the d -dimensional simple symmetric random walk. Also, as it follows from our general bounds (see Proposition 4 and Corollary 7) that the asymptotic results for the genuinely d -dimensional random walk can be reproduced by those of the symmetric *one-dimensional* random walk with appropriately chosen heavy tails, as was indicated by Kesten and Spitzer [2]. The proofs are based on adapting the Tauberian approach developed in [13].

Theorem 3. *Let $d = 1, 2, 3$, and suppose that for $t \in \Gamma := [-\pi, \pi]^d$ one has*

$$\begin{aligned} f(t) &= 1 - \gamma |t| + R(t), \quad \text{for } d = 1, \\ \text{or } f(t) &= 1 - \langle \Sigma t, t \rangle + R(t), \quad \text{for } d = 2, 3, \end{aligned} \quad (6)$$

where Σ is a nonsingular covariance matrix and $R(t) = o(|t|)$ for $d = 1$ and $o(|t|^2)$ for $d = 2, 3$ as $t \rightarrow 0$. Then

$$\begin{aligned} &\text{var}(L_n(\alpha)) \\ &\sim \begin{cases} \frac{(\pi^2 + 6)}{12} \frac{(\alpha!)^2 (\alpha - 1)^2}{(\gamma\pi)^{2\alpha-2}} n^2 \log(n)^{2\alpha-4}, & \text{for } d = 1, \\ \frac{(\alpha!)^2 (\alpha - 1)^2}{2(2\pi\sqrt{|\Sigma|})^{2\alpha-2}} n^2 \log(n)^{2\alpha-4} (\kappa + 1), & \text{for } d = 2, \\ (\kappa_1 + \kappa_2) n \log n, & \text{for } d = 3, \alpha = 2, \end{cases} \end{aligned} \quad (7)$$

where

$$\begin{aligned} \kappa &:= \iint_0^\infty dr ds \left[(1+r)(1+s) \sqrt{(1+r+s)^2 - 4rs} \right]^{-1} \\ &\quad - \frac{\pi^2}{6}, \end{aligned} \quad (8)$$

and κ_1 and κ_2 are defined in (58) and (63), respectively.

Moreover, if $L'(n, \alpha)$ is the self-intersection local time of another random walk, independent of $(S_n)_n$, whose characteristic function also satisfies (6), then $\text{var}(L'_n(\alpha)) = \text{var}(L_n(\alpha))(1 + o(1))$.

2. Proofs

2.1. General Bounds. We first develop a technique to treat random walks with independent but not necessarily identically distributed increments.

Proposition 4 (general upper bound). *Assume that X_1, X_2, \dots are independent \mathbb{Z}^d -valued random variables and let $S_{u,v} := X_u + \dots + X_{u+v}$. Suppose further that for all $n \in \mathbb{N}$ and integers $a, u, b, v \geq 0$, with $a + u \leq b$ and any $x \in \mathbb{Z}^d$, one has*

$$\mathbb{P}(S_{a,u} \pm S_{b,v} = x) \leq \phi(u + v), \quad (\text{A})$$

$$\mathbb{P}(S_{a,u} = 0) - \mathbb{P}(S_{a,u} + S_{b,v} = 0) \leq \psi(u, v), \quad (\text{B})$$

where $\phi(u)$ is nonincreasing and $\psi(u, v)$ is nonincreasing in u and is nondecreasing and subadditive in v in the sense that $\psi(u, v + w) \leq A_\psi[\psi(u, v) + \psi(u, w)]$, for some constant A_ψ independent of u, v , and w . Then, for some constant $K = cA_\psi(1 + A_\psi)^{\alpha-2}$ depending only on α

$$\begin{aligned} \text{var}(L_n(\alpha)) &\leq Kn \left(\sum_{i=0}^{n-1} \phi(i) \right)^{2\alpha-4} \\ &\cdot \sum_{i,j,k=0}^{n-1} [\phi(j \vee i) \phi(k \vee i) + \phi(j) \psi(i + k, j)]. \end{aligned} \quad (9)$$

Proof of Proposition 4. We first write out the variance as a sum

$$\text{var}(L_n(\alpha)) = (\alpha!)^2$$

$$\begin{aligned} I_n &:= \sum_{\substack{k_1 \leq \dots \leq k_\alpha \\ l_1 \leq \dots \leq l_\alpha \\ k_1 \leq l_1, \nu(\delta) \geq 3}} \mathbb{P}[S_{k_1} = \dots = S_{k_\alpha}, S_{l_1} = \dots = S_{l_\alpha}] \\ &= \sum_{x, y \in \mathbb{Z}^d} \sum_{p_1 \leq \dots \leq p_{2\alpha} \leq n} \sum_{\epsilon: \nu(\delta) \geq 3} \mathbb{P}[S_{p_1} = x, S_{p_2} = x + \epsilon_2 y, \dots, S_{p_{2\alpha}} = x + \epsilon_{2\alpha} y] \\ &\leq \sum_{x, y \in \mathbb{Z}^d} \sum_{m_0, \dots, m_{2\alpha-1} \leq n} \sum_{\delta: \nu(\delta) \geq 3} \mathbb{P}(S_{m_0} = x) \mathbb{P}(S_{m_0, m_1} = \delta_1 y) \dots \mathbb{P}(S_{m_{2\alpha-2}, m_{2\alpha-1}} = \delta_{2\alpha-1} y) \\ &= \sum_{y \in \mathbb{Z}^d} \sum_{m_0, \dots, m_{2\alpha-1} \leq n} \sum_{\delta: \nu(\delta) \geq 3} \mathbb{P}(S_{m_0, m_1} = \delta_1 y) \dots \mathbb{P}(S_{m_{2\alpha-2}, m_{2\alpha-1}} = \delta_{2\alpha-1} y). \end{aligned} \quad (11)$$

Summing over the free index m_0 , it is clear that

$$\begin{aligned} I_n &\leq (n + 1) \\ &\cdot \sum_{m_1, \dots, m_{2\alpha-1}} \sum_{y \in \mathbb{Z}^d} \sum_{\delta: \nu(\delta) \geq 3} \prod_{t=1}^{2\alpha-1} \sup_w \mathbb{P}(S_{w, m_t} = \delta_t y). \end{aligned} \quad (12)$$

For any $\delta = (\delta_1, \dots, \delta_{2\alpha-1})$ with $\nu(\delta) = \nu$, exactly $u := 2\alpha - 1 - \nu$ elements are equal to 0, and therefore by Assumption (A) with $x = 0$ we have

$$\begin{aligned} I_n &\leq C(n + 1) \sum_{\nu=3}^{\alpha} \left[\sum_{i=0}^n \phi(i) \right]^{2\alpha-1-\nu} \\ &\cdot \sum_{j_1, \dots, j_\nu=0}^n \sum_{y \in \mathbb{Z}^d} \sum_{\delta' \in \{-1, +1\}^\nu} \prod_{t=1}^{\nu} \sup_{w_t} \mathbb{P}(S_{w_t, j_t} = \delta'_t y). \end{aligned} \quad (13)$$

$$\begin{aligned} &\cdot \sum_{k_1 \leq \dots \leq k_\alpha} \sum_{l_1 \leq \dots \leq l_\alpha} (\mathbb{P}[S_{k_1} = \dots = S_{k_\alpha}, S_{l_1} = \dots = S_{l_\alpha}] \\ &- \mathbb{P}[S_{k_1} = \dots = S_{k_\alpha}] \mathbb{P}[S_{l_1} = \dots = S_{l_\alpha}]). \end{aligned} \quad (10)$$

An important role is played by the manner in which the two sequences are interlaced, since, for example, if $k_\alpha \leq l_1$ or $l_\alpha \leq k_1$, the term vanishes by the Markov property.

We will treat the sum over indices with $k_1 \leq l_1$. The sum over the remaining index set with $k_1 > l_1$ can be treated in a similar fashion and will contribute a constant factor. Therefore, we assume that $k_1 \leq l_1$ and we arrange the two sequences in an ordered sequence of combined length 2α which we denote as $(p_1, \dots, p_{2\alpha})$; we also define $(\epsilon_1, \dots, \epsilon_{2\alpha})$ where $\epsilon_i = 0$ if p_i came from $\mathbf{k} := \{k_1, \dots, k_\alpha\}$ and $\epsilon_i = 1$ if p_i came from $\mathbf{l} := \{l_1, \dots, l_\alpha\}$. Finally we define two new sequences $m_0, m_1, \dots, m_{2\alpha-1}$, and $\delta_1, \dots, \delta_{2\alpha-1}$, where $m_0 := p_1$, $m_i = p_{i+1} - p_i$, and $\delta_i = \epsilon_{i+1} - \epsilon_i$, for $i = 1, \dots, 2\alpha - 1$. Notice that since we assume that $k_1 \leq l_1$, we have $p_1 = k_1$ and $\epsilon_1 = 0$. Let $\nu(\delta) := \sum_{i=1}^{2\alpha-1} |\delta_i|$ denote the *interlacement index*. The terms with $\nu = 1$ vanish, while the terms with $\nu = 2$ will be considered separately.

Terms with $\nu \geq 3$. We first consider the sum I_n over the terms with $\nu \geq 3$ for which we drop the negative part and obtain the bound

Letting $(\tilde{S}_n)_{n \in \mathbb{N}_0}$ denote an independent copy of the random walk $(S_n)_{n \in \mathbb{N}_0}$ and assuming without loss of generality that $j_1 \leq \dots \leq j_\nu$, we have that for any $\delta \in \{-1, +1\}^\nu$

$$\begin{aligned} &\sum_{y \in \mathbb{Z}^d} \prod_{t=1}^{\nu} \sup_{w_t} \mathbb{P}(S_{w_t, j_t} = \delta_t y) \\ &\leq \left(\prod_{t=2}^{\nu-1} \sup_y \sup_{w_t} \mathbb{P}(S_{w_t, j_t} = y) \right) \\ &\cdot \sup_{w_1, w_\nu} \mathbb{P}(S_{w_1, j_1} - \delta_\nu \tilde{S}_{w_\nu, j_\nu} = 0) \leq \left[\prod_{t=2}^{\nu-1} \phi(j_t) \right] \\ &\cdot \phi(j_1 + j_\nu) \leq \prod_{t=2}^{\nu} \phi(j_t \vee j_1). \end{aligned} \quad (14)$$

Let $G_n := \sum_{i=0}^n \phi(i)$. Since ϕ is nonincreasing we have that

$$\begin{aligned} \Delta_{n,\nu} &:= \sum_{0 \leq j_1 \leq \dots \leq j_\nu \leq n} \prod_{t=2}^{\nu} \phi(j_t \vee j_1) \\ &\leq \sum_{j_\nu=0}^n \phi(j_\nu) \sum_{0 \leq j_1 \leq \dots \leq j_{\nu-1} \leq n} \prod_{t=2}^{\nu-1} \phi(j_t \vee j_1) \\ &= G_n \Delta_{n,\nu-1}, \end{aligned} \quad (15)$$

and iterating this procedure, for $\nu \geq 3$, we have that $\Delta_{n,\nu} \leq \Delta_{n,3} G_n^{\nu-3}$. Combining the two bounds and summing over $\nu = 3, \dots, 2\alpha - 1$, we have that

$$\begin{aligned} I_n &\leq \sum_{\nu=3}^{2\alpha-1} c(\alpha) n G_n^{2\alpha-1-\nu} \Delta_{n,\nu} \leq c(\alpha) n G_n^{2\alpha-1-\nu+3} \Delta_{n,3} \\ &= c(\alpha) n G_n^{2\alpha-4} \Delta_{n,3}, \end{aligned} \quad (16)$$

where $c(\alpha)$ is a constant depending only on α .

Terms with $\nu = 2$. Next we consider the sum J_n over the terms with $\nu = 2$, which occurs when, for some j , the indices l_1, \dots, l_α all lie in $[k_j, k_{j+1}]$. Then it is easy to see that this sum J_n is bounded above by

$$\begin{aligned} J_n &\leq Cn \sup_{w_0, \dots, w_{2\alpha-1}} \sum_{m_{\alpha+1}, \dots, m_{2\alpha-2}=0}^n \prod_{r=\alpha+1}^{2\alpha-2} \mathbb{P}(S_{w_r, m_r} = 0) \\ &\cdot \sum_{m_0, \dots, m_\alpha=0}^n \left[\prod_{t=1}^{\alpha-1} \mathbb{P}(S_{w_t, m_t} = 0) \right] \left[\mathbb{P}(S_{w_0, m_0} + S_{w_\alpha, m_\alpha} \right. \\ &= 0) - \mathbb{P}(S_{w_0, m_0} + \dots + S_{w_\alpha, m_\alpha} = 0) \Big] \leq Cn G_n^{\alpha-2} \\ &\cdot \sup_{w_0, \dots, w_\alpha} \sum_{m_0, \dots, m_\alpha=0}^n \left[\prod_{t=1}^{\alpha-1} \mathbb{P}(S_{w_t, m_t} = 0) \right] \\ &\cdot \left[\mathbb{P}(S_{w_0, m_0} + S_{w_\alpha, m_\alpha} = 0) \right. \\ &\left. - \mathbb{P}(S_{w_0, m_0} + \dots + S_{w_\alpha, m_\alpha} = 0) \right] \\ &\leq Cn G_n^{\alpha-2} \sum_{m_0, \dots, m_\alpha=0}^n \left[\prod_{t=1}^{\alpha-1} \phi(m_t) \right] \psi(m_0 + m_\alpha, m_1 \\ &+ \dots + m_{\alpha-1}) \leq C\alpha n G_n^{\alpha-2} A_\psi (1 + A_\psi)^{\alpha-2} \\ &\cdot \left(\sum_{m_2, \dots, m_{\alpha-1}} \prod_{t=2}^{\alpha-1} \phi(m_t) \right) \sum_{m_0, m_1, m_\alpha} \phi(m_1) \psi(m_0 + m_\alpha, \\ &m_1) \leq C\alpha A_\psi (1 + A_\psi)^{\alpha-2} n G_n^{2\alpha-4} \sum_{i,j,k=0}^n \phi(j) \psi(i \\ &+ k, j). \end{aligned} \quad (17)$$

□

The following corollary provides explicit bounds in the cases that are usually considered in the literature.

Corollary 5. *Assume that the conditions of Proposition 4 are satisfied with $\phi(m) = Tm^{-r}$ and $\psi(m, k) = Tm^{-r-1}(k \wedge m)$. Then,*

$$\begin{aligned} \text{var}(L_n(\alpha)) &\leq c_\alpha T^{2\alpha-2} \begin{cases} n^2 \log(n)^{2\alpha-4}, & \text{if } r = 1, \\ n^{4-2r}, & \text{if } 1 < r < \frac{3}{2}, \\ n \log(n), & \text{if } r = \frac{3}{2}, \\ n, & \text{if } r > \frac{3}{2}. \end{cases} \end{aligned} \quad (18)$$

It is straightforward to see that Corollary 5 includes random walks with mean zero and finite second moment; for example, $d = 2$ corresponds to $r = 1$ and $d = 3$ to $r = 3/2$. Therefore several relevant results in [3, 7–13] are obtained as a special case of Corollary 5 and extended to the case of independent but not necessarily identically distributed variables, for example, by applying the local limit theorem, as conducted in [8].

Also when the random walk increment X is in the domain of attraction of the one-dimensional symmetric Cauchy law [13, 14] or in the case of planar random walk with second moments [3, 7–9, 11], it is well known that the conditions of Proposition 4 are satisfied with $\phi(m) = T/m$ and $\psi(m, k) = Tm^{-2}(k \wedge m)$.

However, we can do better for symmetric variables and show that condition (A) implies (B), which together with the comparison technique motivates the following results. For a real number x , we write $[x]$ for the integer part of x .

Proposition 6 (bounds via comparison with characteristic function of symmetric random variables). *Let X_1, X_2, \dots be independent \mathbb{Z}^d -valued random variables and let $f_i(t) := \mathbb{E} \exp(itX_i)$. Assume that there exist a measurable function $f: \Gamma \rightarrow [0, 1]$ and a positive nonincreasing sequence $(\phi(m))_{m \in \mathbb{N}_0}$, such that*

$$\begin{aligned} |1 - f_i(t)| &\leq Tf(t), \\ |f_i(\pm t)| &\leq f(t), \end{aligned} \quad (19)$$

$$\int_\Gamma f(t)^m dt \leq \phi(m),$$

for all integers $i, m \geq 0$, all $t \in \Gamma$, and some positive constant T . Then there exists another positive constant $K = c(\alpha, d, T)$ such that

$$\begin{aligned} \text{var}(L_n(\alpha)) &\leq Kn \left(\sum_{i=0}^{n-1} \phi\left(\left[\frac{i}{2}\right]\right) \right)^{2\alpha-4} \sum_{j=0}^n j \phi\left(\left[\frac{j}{2}\right]\right) \sum_{k=j}^{2n} \phi\left(\left[\frac{k}{2}\right]\right). \end{aligned} \quad (20)$$

Proof of Proposition 6. Using the notation of Proposition 4, for positive integers a, u, b , and v , with $a + u \leq b$, $\epsilon_j = \pm 1$, and any $x \in \mathbb{Z}^d$

$$\begin{aligned} & \mathbb{P}(S_{a,u} + \epsilon \cdot S_{b,v} = x) \\ & \leq \frac{1}{(2\pi)^d} \int_{\Gamma} \prod_{j \in [a,a+u] \cup [b,b+v]} |f_j(\epsilon_j t)| dt \\ & \leq \frac{1}{(2\pi)^d} \int_{\Gamma} f(t)^{u+v} dt \leq \frac{1}{(2\pi)^d} \phi(u+v). \end{aligned} \tag{21}$$

To find $\psi(u, v)$, notice that since $f(t) \geq 0$,

$$\begin{aligned} \phi(m) & \geq \int_{\Gamma} f(t)^m [1 - f(t)^m] dt \\ & = \sum_{j=0}^{m-1} \int_{\Gamma} f(t)^{m+j} (1 - f(t)) dt \\ & \geq m \int_{\Gamma} f(t)^{2m} (1 - f(t)) dt =: mQ(2m) \end{aligned} \tag{22}$$

whence $Q(m) \leq 2\phi([m/2])/m$. Therefore,

$$\begin{aligned} & |\mathbb{P}(S_{a,u} = 0) - \mathbb{P}(S_{a,u} + S_{b,1} = 0)| \\ & = \left| \frac{1}{(2\pi)^d} \int_{\Gamma} \left[\prod_{j=a}^{a+u} f_j(t) \right] (1 - f_{b+1}(t)) dt \right| \\ & \leq CT \int_{\Gamma} |f(t)|^u |1 - f(t)| dt \leq \frac{CT\phi([u/2])}{u}. \end{aligned} \tag{23}$$

A telescoping argument implies that

$$|\mathbb{P}(S_{a,u} = 0) - \mathbb{P}(S_{a,u} + S_{b,v} = 0)| \leq CT\phi\left(\left[\frac{u}{2}\right]\right) \frac{v}{u}. \tag{24}$$

On the other hand for $u \leq v$ we can obtain a tighter bound through

$$\begin{aligned} \mathbb{P}(S_{a,u} = 0) - \mathbb{P}(S_{a,u} + S_{b,v} = 0) & \leq \mathbb{P}(S_{a,u} = 0) \\ & \leq \phi(u). \end{aligned} \tag{25}$$

Combining the two bounds above it follows that (B) is satisfied with $\psi(u, v) := \phi([u/2]) \min(u, v)/u$. Thus all conditions of Proposition 4 are satisfied and the result follows. \square

The following corollary allows for the case where $\phi(m)$ is regularly varying.

Corollary 7. Assume that the conditions of Proposition 6 are satisfied with $\phi(m) = h(m)m^{-r}$, $r \geq 1$, where $h(\cdot)$ is slowly varying at ∞ . Then,

$$\begin{aligned} \text{var}(L_n(\alpha)) & \leq K\Delta_n(\alpha, \phi) \\ & \leq c_{\alpha} T^{2\alpha-2} \begin{cases} n^2 \left[\sum_{k=1}^n \frac{h(k)}{k} \right]^{2\alpha-4}, & \text{for } r = 1, \\ n^{4-2r} h^2(n), & \text{for } 1 < r < \frac{3}{2}, \\ n \sum_{k=1}^n \frac{h(k)^2}{k}, & \text{for } r = \frac{3}{2}, \\ n, & \text{for } r > \frac{3}{2}. \end{cases} \end{aligned} \tag{26}$$

Several results in [3, 7–13] are obtained as a special case of Corollary 7 and can be extended to dependent variables, for example, a random walk driven by a hidden Markov chain. In addition, following [2], we can construct a one-dimensional symmetric random walk with characteristic function $f(t) = 1 - c|t|^{1/r} + o(|t|^{1/r})$, where $r = 2/d$ for $d = 2, 3$ and $r = 1/2$ for $d \geq 4$, whose asymptotic behaviour is similar to that of genuinely d -dimensional random walk.

The following example of genuinely 2-dimensional recurrent walk with infinite variance was motivated by Spitzer [1, pp. 87].

Example 8. Let X_1, X_2, \dots be independent, identically distributed, \mathbb{Z}^2 -valued random variables, such that $\mathbb{P}(|X_1| = k) = c/(k^3 \log(k)^g)$, for $k \geq 4$ and $g \in [0, 1)$. Let $(S_n)_{n \in \mathbb{N}_0}$ be the corresponding random walk in \mathbb{Z}^2 . Then we have

$$\begin{aligned} \text{var}(L_n(\alpha)) & \leq cn^2 \max\{[\log n]^g, \log \log n\}^{2\alpha-4} \log n^{-2(1-g)}, \end{aligned} \tag{27}$$

for $n \geq 10$. Under these assumptions we have that $\mathbb{P}(S_n = 0) \leq c/n \log(n)^{1-g}$, which is in the *critical range*, where the random walk is recurrent, without second moment. To see why, we note that by a lengthy but straightforward calculation it can be shown that the characteristic function of X satisfies (19) with

$$\begin{aligned} \phi(n) & = \frac{c}{n \log(e \vee n)^{1-g}}, \\ f(t) & = \exp\left[-A |t|^2 h(|t|^2)\right], \end{aligned} \tag{28}$$

$$\text{where } h(r) := \left[1 + \log\left(\frac{1}{r}\right)_+\right]^{1-g}.$$

The sequence $\phi(m)$ is identified via Fourier inversion, polar coordinates, and a Laplace argument,

$$\begin{aligned} \int_{\Gamma} f(t)^n dt & \leq c \int_0^1 \exp\left(-nr \left(1 + \log\left(\frac{1}{r}\right)\right)^{1-g}\right) \\ & + O(e^{-n}) \leq \frac{c}{n \log(e \vee n)^{1-g}} =: \phi(n). \end{aligned} \tag{29}$$

2.2. Bounds for Identically Distributed Variables

Proposition 9 (general upper bound for i.i.d.). Let X, X_1, X_2, \dots be independent, identically distributed,

\mathbb{Z}^d -valued random variables. Suppose that for any $x \in \mathbb{Z}^d$ and all positive integers a, u, b , and v , with $a + u \leq b$, it holds that

$$\mathbb{P}(S_{a,u} \pm S_{b,v} = x) \leq \phi(u + v), \quad (30)$$

where $\{\phi(m)\}_{m \in \mathbb{N}_0}$ is a nonincreasing sequence. Then for some constant $K = c(\alpha)$ we have that

$$\begin{aligned} & \text{var}(L_n(\alpha)) \\ & \leq Kn \left(\sum_{i=0}^{n-1} \phi(i) \right)^{2\alpha-4} \sum_{j=0}^n j \phi(j) \sum_{k=j}^{[\alpha n]+1} \phi\left(\left\lfloor \frac{k}{\alpha} \right\rfloor\right). \end{aligned} \quad (31)$$

Proof of Proposition 9. By inspecting the proof of Proposition 6, we notice that we only need to bound the term J_n . Consider typical ordering

$$0 \leq i_1 \leq \dots \leq i_k \leq j_1 \leq \dots \leq j_\alpha \leq i_{k+1} \leq \dots \leq i_\alpha \leq n, \quad (32)$$

and let us change variables to $(m_0, \dots, m_{2\alpha})$ such that $m_0 + \dots + m_{2\alpha} = n$. Then the contribution to J_n is given by

$$\begin{aligned} & \sum_{m_0, \dots, m_{2\alpha}} \prod_{\substack{j \neq k, k+\alpha \\ 1 \leq j \leq 2\alpha-1}} \mathbb{P}(S_{m_j} = 0) \\ & \cdot \left[\mathbb{P}(S_{m_k+m_{k+\alpha}} = 0) - \mathbb{P}(S_{m_k+\dots+m_{k+\alpha}} = 0) \right]. \end{aligned} \quad (33)$$

We keep m_j fixed for $j \neq \alpha, k + \alpha$ and we sum over $m = m_k + m_{k+\alpha}$ from 0 to some $M = M(n, \{m_j\}_{j \neq k, k+\alpha})$. Then for given $m_{k+1}, \dots, m_{k+\alpha-1}$, the term in the sum is

$$\sum_{m=0}^M (m+1) \left[\mathbb{P}(S_m = 0) - \mathbb{P}(S_{m+q} = 0) \right], \quad (34)$$

where $q := m_{k+1} + \dots + m_{k+\alpha-1}$. Then since $M \leq n - q$, it is an easy exercise to show that this sum is bounded above by

$$\begin{aligned} & \sum_{m=0}^M (m+1) \left[\mathbb{P}(S_m = 0) - \mathbb{P}(S_{m+q} = 0) \right] \\ & \leq \sum_{m=0}^{q-1} (m+1) \mathbb{P}(S_m = 0) + q \mathbb{1}(n - q \geq q) \\ & \cdot \sum_{m=q}^{n-q} \mathbb{P}(S_m = 0) \leq \sum_{m=0}^{(\alpha m^*) \wedge n} (m+1) \mathbb{P}(S_m = 0) \\ & + \alpha m^* \sum_{m=m^*}^n \mathbb{P}(S_m = 0), \end{aligned} \quad (35)$$

where $m^* = \max\{m_{k+1}, \dots, m_{k+\alpha-1}\}$. The result follows by summing over all indices apart from m^* and changing the order of summation. \square

2.3. Proofs of Main Results

Proof of Theorem 1. We apply a comparison argument found to be useful in many areas (e.g., Montgomery-Smith and Pruss [15], and Lefèvre and Utev [16]). More specifically we

bound the quantity $\text{var}(L_n)$ by the corresponding quantity for the symmetrised random walk.

Following Spitzer's argument we notice that with $f(t) = \mathbb{E}[\exp(it \cdot X_1)]$

$$\begin{aligned} & \mathbb{P}(S_{a,u} + \epsilon S_{b,v} = x) \leq c \int_{\Gamma} |f(t)|^u |f(-t)|^v dt \\ & = c \int_{\Gamma} [|f(t)|^2]^{u/2} [|f(-t)|^2]^{v/2} dt. \end{aligned} \quad (36)$$

Since $|f(t)|^2$ is the characteristic function of a symmetric random variable in \mathbb{Z}^d , for some positive λ , we have $1 - |f(t)|^2 \geq \lambda|t|^2$, and, hence,

$$\begin{aligned} & \mathbb{P}(S_{a,u} + \epsilon S_{b,v} = x) \leq c \int_{\Gamma} \exp\left[-\frac{\lambda(u+v)}{2}|t|^2\right] dt \\ & \leq c(u+v)^{-d/2}. \end{aligned} \quad (37)$$

The result follows from Proposition 9 applied with $\phi(m) = m^{-d/2}$. \square

The proof of Theorem 2 will be based on the following lemma.

Lemma 10. *Assume X, X_1, X_2, \dots are independent, identically distributed, genuinely d -dimensional random variables such that $\mathbb{E}|X|^2 = \infty$. Then there exists a monotone, slowly varying sequence $(h_n)_{n \in \mathbb{N}_0}$, such that $h_n \rightarrow 0$ as $n \rightarrow \infty$ and*

$$\sup_{x \in \mathbb{Z}^d} \mathbb{P}(S_n = x) \leq c_d \int_{\Gamma} |\mathbb{E}e^{it \cdot X}|^n dt \leq h_n n^{-d/2}. \quad (38)$$

Proof of Lemma 10. Without loss of generality we assume that X is symmetric. Let $\sigma_{e,L} := \mathbb{E}[(e \cdot X)^2 \mathbb{1}(|X| \leq L)]$. Following Spitzer, since X is genuinely d -dimensional, we may assume that there exist positive constants c, W , such that for any unit vector $|e| = 1$ we have that $\sigma_{e,W} \geq c$ and $1 - f(t) \geq c|t|^2$ for all $t \in \Gamma$. Let λ_d be the d -dimensional Lebesgue measure on \mathbb{R}^d and μ_d the Lebesgue-Haar measure on $S^{d-1} := \{e \in \Gamma : |e| = 1\}$. Notice that since $\mathbb{E}|X|^2 = \infty$, for any K , we have $\mu_d\{e : \sigma_{e,\infty} < K\} = 0$.

Fix a small positive x such that $\sqrt{c/x} \geq 2W$, and for any $\epsilon > 0$ let $K = K(\epsilon) = \epsilon^{-d/2}$. Then there exists $L = L(\epsilon) > 0$ small enough so that $\mu_d\{e : \sigma_{e,L} < K\} \leq \epsilon^{d/2}$. We partition S^{d-1} in two sets

$$\begin{aligned} A_{L,K} &= \{e \in S_d : \sigma_{e,L} \geq K\}, \\ \bar{A}_{L,K} &= \{e \in S_d : \sigma_{e,L} < K\}, \end{aligned} \quad (39)$$

so that, for any direction $e \in \bar{A}_{L,K}$,

$$\begin{aligned} & \{z \in \mathbb{R} : 1 - f(ze) \leq x\} \subseteq \{z : cz^2 \leq x\} \\ & \subseteq \left\{z : |z| \leq \sqrt{\frac{x}{c}}\right\}. \end{aligned} \quad (40)$$

Hence, using d -dimensional spherical coordinates,

$$\begin{aligned} \lambda_d \{ (z, e) \in \mathbb{R} \times \bar{A}_{L,K} : 1 - f(ez) \leq x \} \\ \leq \mu_d \{ \bar{A}_{L,K} \} \left(\frac{x}{c} \right)^{d/2} \left(\frac{1}{d} \right) \leq \epsilon^{d/2} \left(\frac{x}{c} \right)^{d/2} \left(\frac{1}{d} \right). \end{aligned} \quad (41)$$

On the other hand, for any t ,

$$\begin{aligned} 1 - f(t) &= 2 \sum_{k \in \mathbb{Z}^d} \sin \left(\frac{[t \cdot k]}{2} \right)^2 P(X = k) \\ &\geq \left(\frac{1}{4} \right) E \left[(t \cdot X)^2 I \left(|t \cdot X| \leq \frac{1}{2} \right) \right] \\ &= \left(\frac{|t|^2}{4} \right) \sigma_{t/|t|, 1/2|t|}. \end{aligned} \quad (42)$$

Now, assume that $\sqrt{c/x} \geq 2L$. Then for any direction $e \in A_{L,K}$, by choice of x and since $\sigma_{e,L}$ is increasing in L , for $cz^2 \leq 1 - f(ez) \leq x$ or $|z| \leq \sqrt{x/c}$, it must be the case that

$$\begin{aligned} x \geq 1 - f(ez) &\geq \left(\frac{z^2}{4} \right) \sigma_{e, 1/2z} \geq \left(\frac{z^2}{4} \right) \sigma_{e,L} \\ &\geq \left(\frac{z^2}{4} \right) K, \end{aligned} \quad (43)$$

implying that, on the set $A_{L,K}$, it must be that $|z| \leq 2\sqrt{x/K}$. Changing to d -dimensional polar coordinates, we find that

$$\begin{aligned} \lambda_d \{ (z, e) \in \mathbb{R} \times A_{L,K} : 1 - f(ez) \leq x \} \\ \leq \int_{A_{L,K}} \int_0^{\sqrt{4x/K}} r^{d-1} dr de \leq C_d \epsilon^{d/2} x^{d/2}. \end{aligned} \quad (44)$$

Overall, for $x \leq c/4L^2$, $\lambda_d \{ t : 1 - f(t) \leq x \} \leq c_d(x\epsilon)^{d/2}$, and hence $\{ t \in \Gamma : 1 - f(t) \leq x \}$ has Lebesgue measure $o(x^{d/2})$.

Let $F(x)$ be the cumulative distribution function of the random variable $\log(1/f(\cdot))$ defined on the probability space Γ with normalised Lebesgue measure. Then F is continuous at $x = 0$ and supported on \mathbb{R}^+ . Moreover, we have that $F(x) = o(x^{d/2})$ as $x \downarrow 0$. Therefore, for some positive sequence $(\epsilon_n)_{n \in \mathbb{N}_0}$ with $\epsilon_n \rightarrow 0$, we have that

$$\begin{aligned} \frac{1}{(2\pi)^2} \int_{\Gamma} f(t)^n dt &= \int_0^{\infty} e^{-nx} dF(x) \\ &= n \int_0^{\infty} e^{-nx} F(x) dx \leq n^{-d/2} \epsilon_n. \end{aligned} \quad (45)$$

It remains to show that there exists a positive, monotone, slowly varying sequence $(h_n)_{n \in \mathbb{N}_0}$, such that $\epsilon_n \leq h(n) \rightarrow 0$ as $n \rightarrow \infty$. Let $\delta_n = \sup_{j \geq n} \epsilon_j$ and $a_0 := 0$ and for $n \geq 1$ define a_n recursively by $a_n = \min(2a_{2^{r-1}}, 1/\delta_n)$, for $2^{r-1} < n \leq 2^r$, so that $a_n \rightarrow \infty$ is monotone, $a_{2^r} \leq 2a_{2^{r-1}}$ implying that $a_{2^n} \leq 4a_n$, and $1/a_n \geq \delta_n \geq \epsilon_n$. Finally, take $h_n := 1/\max(a_0, \log a_n)$. \square

Proof of Theorem 2. Assume that $\mathbb{E}|X|^2 = \infty$ and $d = 2$ or $d = 3$. Then, by Lemma 10 there exists a slowly varying sequence $h_n \rightarrow 0$ as $n \rightarrow \infty$ such that $\int_{\Gamma} |\mathbb{E} \exp(it \cdot X)|^n dt \leq h_n n^{-d/2}$. Applying Corollary 7 with $r = 1$ and $r = 3/2$ we, respectively, find that

$$\begin{aligned} \text{var}(L_n(\alpha)) \\ \leq \begin{cases} Kn^2 \left(\sum_{k=1}^n \frac{h(k)}{k} \right)^{2\alpha-4} = o(n^2 (\log n)^{2\alpha-4}), & \text{for } d = 2, \\ Kn \left(\sum_{k=1}^n \frac{h(k)^2}{k} \right) = o(n \ln n), & \text{for } d = 3. \end{cases} \end{aligned} \quad (46)$$

Finally assume that $\mathbb{E}|X|^2 < \infty$ and $E[X] = \mu \neq 0$. Then $\mathbb{P}(S_n = 0) = \mathbb{P}(S'_n = -n\mu)$ whence it follows that $\mathbb{P}(S_n = 0) = o(n^{-d/2})$ (see, e.g., [17, Theorem 2.3.10]). Then inspecting the proof of Proposition 4, one can readily obtain the desired bound for the J_n term, while with slight modification the bound for the I_n term also follows.

Note that for $d = 1$ the situation is much simpler since then $\text{var}(L_n^{\text{SRW}}(\alpha)) \sim C[\mathbb{E}L_n^{\text{SRW}}(\alpha, d)]^2$ and if $\mathbb{E}|X|^2 = \infty$ or $\mathbb{E}[X] \neq 0$, $\mathbb{E}L_n^{\text{SRW}}(\alpha, d) = o(n^{(1+\alpha)/2})$. \square

Proof of Theorem 3. We first give the proof for the case $d = 1$. As in the proof of Proposition 4 we begin from expression (10) and define the sequences p_i and δ_i for $i = 1, \dots, 2\alpha - 1$, and the quantity $\nu(\delta) = \sum_{i=1}^{2\alpha-1} |\delta_i|$. Recall that $\nu(\delta)$ measures the interlacement of the two sequences k_1, \dots, k_α and l_1, \dots, l_α . For example, $\nu(\delta) = 1$ occurs when either $k_\alpha \leq l_1$ or $l_\alpha \leq k_1$, in which case the contribution vanishes by the Markov property. On the other hand $\nu(\delta) = 2$ when, for example, $l_1, \dots, l_\alpha \in [k_i, k_{i+1}]$ for some i . Finally $\nu(\delta) = 3$ occurs when, for example,

$$\begin{aligned} k_1 \leq \dots \leq k_r \leq l_1 \leq \dots \leq l_s \leq k_{r+1} \leq \dots \leq k_\alpha \leq l_{s+1} \\ \leq \dots \leq l_\alpha \leq n. \end{aligned} \quad (47)$$

From the proof of Proposition 4, and using the bound $\mathbb{P}(S_n = 0) \leq c/n$, the terms of the sum are bounded above by $n^2 \log(n)^{2\alpha-1-\nu(\delta)}$, and thus the leading term appears when either $\nu(\delta) = 2, 3$, with other terms giving strictly lower order. We will therefore analyze these two situations in detail in order to derive the exact asymptotic constants. When $\nu = 3$, the two terms in the difference individually give the correct order and will be treated by the classical Tauberian theory. However for $\nu = 2$, the two terms only give the correct order when considered together. This however forbids the use of Karamata's Tauberian theorem since the monotonicity restriction would require roughly that X_i is symmetric. Thus the complex Tauberian approach, as developed in [13], is required to justify the answer.

Case 1 ($\nu(\delta) = 3$). Assume that part of the sequence $\mathbf{l} = \{l_1, \dots, l_\alpha\}$ lies between k_r and k_{r+1} and the rest between k_s and k_{s+1} . Then using the change of variables

$$\begin{aligned}
i_1 &= m_0, \\
i_2 &= m_0 + m_1, \\
&\vdots \\
i_r &= m_0 + \cdots + m_{r-1} \\
j_1 &= m_0 + \cdots + m_r, \\
j_2 &= m_0 + \cdots + m_{r+1}, \\
&\vdots \\
j_s &= m_0 + \cdots + m_{r+s-1}, \\
i_{r+1} &= m_0 + \cdots + m_{r+s}, \\
i_{r+2} &= m_0 + \cdots + m_{r+s+1}, \\
&\vdots \\
i_\alpha &= m_0 + \cdots + m_{\alpha+s-1}, \\
j_{s+1} &= m_0 + \cdots + m_{\alpha+s}, \\
j_{s+2} &= m_0 + \cdots + m_{\alpha+s+1}, \\
&\vdots \\
j_\alpha &= m_{2\alpha-1}, \\
n &= m_0 + \cdots + m_{2\alpha},
\end{aligned} \tag{48}$$

we rewrite the positive term in (10) as

$$\begin{aligned}
a(n) &= \sum \mathbb{P} [S(i_1) = \cdots = S(i_\alpha); S(j_1) = \cdots = S(j_\alpha)] \\
&= \sum_{m_0, \dots, m_{2\alpha-1}} \left[\prod_{\substack{j=1 \\ j \neq r, r+s, \alpha+s}}^{2\alpha-1} \mathbb{P}(S_{m_j} = 0) \right] \\
&\cdot \mathbb{P}(S_{m_r} + S'_{m_{r+s}} = S'_{m_{r+s}} + S''_{m_{\alpha+s}} = 0).
\end{aligned} \tag{49}$$

Notice that from [13] we have that $\sum_{n \geq 0} \lambda^n \mathbb{P}(S_n = 0) \sim \log(1/(1-\lambda))/\pi\gamma$. Let

$$\begin{aligned}
a(\lambda) &= (1-\lambda)^{-3} [-\log(1-\lambda)]^{2\alpha-4}, \\
c_\gamma &= (\pi\gamma)^{-2\alpha+4}.
\end{aligned} \tag{50}$$

Then, by direct calculations and Fourier inversion formula

$$\begin{aligned}
\sum_{n \geq 0} \lambda^n a(n) &= c_\gamma (1-\lambda) a(\lambda) \\
&\cdot \sum_{x \in \mathbb{Z}} \sum_{k_1, k_2, k_3 \geq 0} \lambda^{k_1+k_2+k_3} \mathbb{P}(S_{k_1} = x) \mathbb{P}(S_{k_2} = -x) \\
&\cdot \mathbb{P}(S_{k_3} = x) = c_\gamma (1-\lambda) a(\lambda) \frac{1}{(2\pi)^2} \\
&\cdot \iint_{\Gamma} \frac{dt ds}{(1-\lambda f(t))(1-\lambda f(s))(1-\lambda f(t+s))} \\
&\sim c_\gamma (1-\lambda) a(\lambda) \frac{1}{(2\pi)^2} \frac{1}{\gamma^2} \frac{1}{1-\lambda} \\
&\cdot \iint_{\mathbb{R}^2} \frac{dx dy}{(1+|x|)(1+|y|)(1+|x+y|)} \sim \left(\frac{1}{4\gamma^2}\right) \\
&\cdot c_\gamma a(\lambda).
\end{aligned} \tag{51}$$

Next we consider the negative term in (10)

$$\begin{aligned}
b(n) &:= \sum_{m_0, \dots, m_{2\alpha-1}} \mathbb{P} [S_{m_1} = \cdots = S_{m_{r-1}} = S_{m_r} + \cdots \\
&+ S_{m_{r+s}} = S_{m_{r+s+1}} = \cdots = S_{m_{\alpha+s-1}} = 0] \mathbb{P} [S_{m_{r+1}} = \cdots \\
&= S_{m_{r+s}} + \cdots + S_{m_{\alpha+s}} = S_{m_{\alpha+s+1}} = \cdots = S_{m_{2\alpha-1}} = 0].
\end{aligned} \tag{52}$$

By direct calculations and (6),

$$\begin{aligned}
\sum_n \lambda^n b(n) &= \left(\frac{1}{\pi\gamma} \log\left(\frac{1}{1-\lambda}\right)\right)^{\alpha-s+r-2} (1-\lambda)^{-2} \\
&\cdot \sum_{m_r, \dots, m_{\alpha+s}=0}^{\infty} \lambda^{m_r + \cdots + m_{\alpha+s}} \\
&\cdot \prod_{\substack{t=r+1, \dots, \alpha+s-1 \\ t \neq r+s}} \mathbb{P}(S_{m_t} = 0) \\
&\cdot \mathbb{P}(S_{m_r} + \cdots + S_{m_{r+s}} = 0) \\
&\cdot \mathbb{P}(S_{m_{r+s}} + \cdots + S_{m_{\alpha+s}} = 0),
\end{aligned} \tag{53}$$

and using Fourier inversion and (6) the internal sum behaves as

$$\begin{aligned}
(2\pi)^{-\alpha-s+r} \int_{-\pi}^{\pi} \cdots \int_{-\pi}^{\pi} (1-\lambda\phi(x))^{-1} (1-\lambda\phi(x)\phi(y))^{-1} (1-\lambda\phi(y))^{-1} \\
\cdot \left[\prod_{j=r+1}^{r+s-1} \prod_{k=r+s+1}^{\alpha+s-1} (1-\lambda\phi(x)\phi(t_j))^{-1} (1-\lambda\phi(y)\phi(t_k))^{-1} dt_j dt_k \right] dx dy \sim (\pi\gamma)^{-\alpha-s+r} (1-\lambda)^{-1} \\
\cdot \log\left(\frac{1}{1-\lambda}\right)^{\alpha-r+s-2} \frac{\pi^2}{6}.
\end{aligned} \tag{54}$$

Then, we have $\sum_n \lambda^n b(n) \sim (\pi^2/6(\pi\gamma)^{2\alpha-2})a(\lambda)$, whence the Tauberian theorem implies that $a(n) - b(n) \sim n^2 \log(n)^{2\alpha-4} / 24\pi^{2\alpha-4} \gamma^{2\alpha-2}$. Most importantly we see that the lengths and locations of the chains, r and s , do not affect the asymptotic behaviour. Noting that if $1 \leq r, s \leq \alpha - 1$, we can partition $2\alpha = r + s + (\alpha - r) + (\alpha - s)$ in $(\alpha - 1)^2$ ways, and thus overall the total contribution from terms with $\nu = 3$ is

$$\left[\frac{(\alpha! (\alpha - 1)^2}{12\pi^{2\alpha-4} \gamma^{2\alpha-2}} \right] n^2 \log(n)^{2\alpha-4}. \tag{55}$$

Case 2 ($\nu(\delta) = 2$). The typical term $c(n)$ was introduced in (33) in the proof of Proposition 9. Now we let $\lambda \in \mathbb{C}$, with $|\lambda| < 1$. By lengthy but direct calculations we can derive an expression of the form

$$\sum_n \lambda^n c(n) = \frac{\alpha - 1}{(\gamma\pi)^{2\alpha-2}} a(\lambda) + o(a(\lambda)), \quad \lambda \rightarrow 1. \tag{56}$$

The approach developed in [13] can then be used to bound the error terms and show that $c(n) \sim [(\alpha - 1)/2(\gamma\pi)^{2\alpha-2}] n^2 \log(n)^{2\alpha-4}$.

Finally taking into account the fact that l_1, \dots, l_α can be in any of the $\alpha - 1$ intervals $[k_i, k_{i+1}]$, for $i = 1, \dots, \alpha - 1$, the result follows the overall contribution of terms with $\nu(\delta) = 2$

$$\frac{(\alpha - 1)^2}{2(\gamma\pi)^{2\alpha-2}} n^2 \log(n)^{2\alpha-4}. \tag{57}$$

The case for $d = 2$ is very similar, so we move on to the case $d = 3$.

Case 3 ($d = 3$ and $\alpha = 2$). Using the same notation as before, we have three terms to consider $a(n), b(n)$, and $c(n)$. We first consider $c(n)$. Letting $K := \epsilon/\sqrt{1 - \lambda}$ and using the usual power series construction and spherical coordinates

$$\sum_n \lambda^n c(n) = (1 - \lambda)^{-2} (2\pi)^{-6}$$

$$\begin{aligned} I_1(\lambda) &\sim |\Sigma|^{-1} \int_{r,s=0}^\epsilon \int_{\phi_{1,2}=0}^{2\pi} \int_{\theta_1,\theta_2=0}^\pi \frac{r^2 s^2 \sin(\theta_1) \sin(\theta_2) d\theta_1 d\theta_2 d\phi_1 d\phi_2 dr ds}{(1 - \lambda + \lambda r^2)(1 - \lambda + \lambda s^2)[1 - \lambda + \lambda(r^2 + s^2 + 2Ars)]} \\ &= |\Sigma|^{-1} \int_{\theta_1,\theta_2=0}^\pi \int_{\phi_1,\phi_2=0}^{2\pi} \int_{r,s=0}^K \frac{\sin(\theta_1) \sin(\theta_2) r^2 s^2 ds dr d\phi_1 d\phi_2 d\theta_1 d\theta_2}{(1 + r^2)(1 + s^2)[1 + r^2 + s^2 + 2Ars]} \\ &\sim |\Sigma|^{-1} \log(K) \int_{\theta_1,\theta_2=0}^\pi \int_{\phi_1,\phi_2=0}^{2\pi} \sin(\theta_1) \sin(\theta_2) \frac{\arccos(A(\theta_1, \theta_2, \phi_1, \phi_2))}{\sqrt{1 - A(\theta_1, \theta_2, \phi_1, \phi_2)^2}} d\phi_1 d\phi_2 d\theta_1 d\theta_2. \end{aligned} \tag{61}$$

The other integral is slightly easier

$$\begin{aligned} I_2(\lambda) &\sim |\Sigma|^{-1} \frac{\pi}{2} \log K \\ &\cdot \int_{\theta_1,\theta_2=0}^\pi \int_{\phi_1,\phi_2=0}^{2\pi} \sin(\theta_1) \sin(\theta_2) d\phi_1 d\phi_2 d\theta_1 d\theta_2, \end{aligned} \tag{62}$$

$$\begin{aligned} &\cdot \iint_{J^3 \times J^3} \frac{\lambda f(y)(1 - f(x)) dx dy}{(1 - \lambda f(x))^2 (1 - \lambda f(y)) (1 - \lambda f(x) f(y))} \\ &\sim 2(2\pi)^{-4} |\Sigma|^{-1} (1 - \lambda)^{-2} \\ &\cdot \iint_0^K \frac{r^4 s^2 dr ds}{(1 + r^2)^2 (1 + s^2)^2 (1 + r^2 + s^2)} \sim 2(2\pi)^{-4} |\Sigma|^{-1} \\ &\cdot \frac{\pi}{2} (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right) =: \kappa_1 (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right), \end{aligned} \tag{58}$$

and thus $c(n) \sim \kappa_1 n \log n$, where $\kappa_1 > 0$, where the answer can be justified following [13].

The term $a(n) - b(n)$ is trickier to compute. As usual we consider the power series

$$\begin{aligned} \sum_{n \geq 0} \lambda^n (a(n) - b(n)) &= (1 - \lambda)^{-2} (2\pi)^{-6} \\ &\cdot \iint_{B(\epsilon)} \frac{dx dy}{(1 - \lambda f(x))(1 - \lambda f(y))(1 - \lambda f(x + y))} \\ &- (1 - \lambda)^{-2} (2\pi)^{-6} \\ &\cdot \iint_{B(\epsilon)} \frac{dx dy}{(1 - \lambda f(x))(1 - \lambda f(y))(1 - \lambda f(x) f(y))} \\ &= (1 - \lambda)^{-2} (2\pi)^{-6} (I_1(\lambda) - I_2(\lambda)). \end{aligned} \tag{59}$$

Let $A \in [-1, 1]$ be the cosine of the angle between x and y , which in spherical coordinates is

$$\begin{aligned} A &= A(\theta_1, \theta_2, \phi_1, \phi_2) \\ &= \cos(\phi_1 - \phi_2) \sin(\theta_1) \sin(\theta_2) \\ &\quad + \cos(\theta_1) \cos(\theta_2). \end{aligned} \tag{60}$$

Then as $0 < \lambda \uparrow 1$, using the expansion (6)

and thus overall we must have that

$$\begin{aligned} (I_1 - I_2)(\lambda) &\sim \frac{1}{2} (2\pi)^{-6} |\Sigma|^{-1} (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right) \\ &\cdot \int_{\theta_1,\theta_2=0}^\pi \int_{\phi_1,\phi_2=0}^{2\pi} \left[\frac{\arccos(A)}{\sqrt{1 - A^2}} - \frac{\pi}{2} \right] \sin(\theta_1) \end{aligned}$$

$$\begin{aligned} & \cdot \sin(\theta_2) d\phi_1 d\phi_2 d\theta_1 d\theta_2 =: \kappa_2 (1 \\ & - \lambda)^{-2} \log\left(\frac{1}{1-\lambda}\right), \end{aligned} \quad (63)$$

whence it follows that $\text{var}(L_n(2)) \sim (\kappa_1 + \kappa_2)n \log n$.

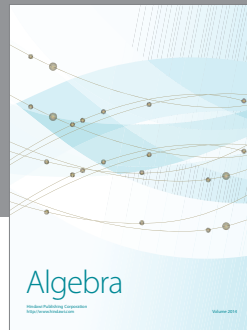
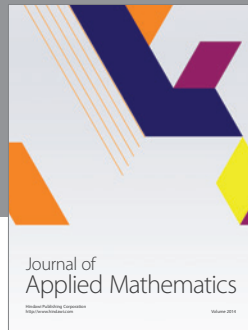
To prove the last claim let $S'_n = X'_1 + \dots + X'_n$ be another random walk, independent of S_n , such that its characteristic function $f'(t) = \mathbb{E}[\exp(itX'_i)]$ also satisfies the expansion (6). Then using [13, Lemma 3.1], one can adapt the proof of [13, Theorem 2.1] to show that $L'_n(\alpha) = L_n(\alpha) + o(L_n(\alpha))$. \square

Competing Interests

The authors declare that they have no competing interests.

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